# Advances in Laplace Inference for Reliable and Interpretable Bayesian Deep Learning

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#### Motivation



Figure: Uncertainty for semantic segmentation. (Fig. 1 of Kendall et al., 2017)

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#### Laplace approximation: Notation

We will be focusing on supervised learning over inputs  $x_{1:n}$  and labels  $y_{1:n}$ .

**Neural network:** Let  $f_{\theta} : \mathcal{X} \to \mathcal{Y}$  be a neural network with parameters  $\theta$ .

$$f_{oldsymbol{ heta}} riangleq f_{oldsymbol{ heta}^{(L)}}^{(L)} \circ \cdots \circ f_{oldsymbol{ heta}^{(1)}}^{(1)}, \ oldsymbol{ heta} riangleq extsf{Concat}ig[oldsymbol{ heta}^{(1)}, \dots, oldsymbol{ heta}^{(L)}ig] \in \mathbb{R}^{P}, \ oldsymbol{ heta}^{(\ell)} \in \mathbb{R}^{D_{ extsf{in}} \cdot D_{ extsf{out}}}$$

**Probabilistic inference:** Update our belief over  $\theta$  after seeing  $\{x_{1:n}, y_{1:n}\}$ .



Variational inference: Goal is to approximate the posterior distribution.

$$p(\theta \mid \mathbf{x}_{1:n}, y_{1:n}) = \frac{p(y_{1:n} \mid \mathbf{x}_{1:n}, \theta) p(\theta)}{\int p(y_{1:n} \mid \mathbf{x}_{1:n}, \theta) p(\theta)} \approx q(\theta \mid \lambda) \triangleq q_{\lambda}(\theta)$$

MAP estimation: Rely on SGD to find a maximum a posteriori estimate.

$$\hat{\boldsymbol{\theta}} = \arg \max_{\boldsymbol{\theta}} p(\boldsymbol{\theta} \mid \boldsymbol{x}_{1:n}, y_{1:n})$$
  
=  $\arg \min_{\boldsymbol{\theta}} - \log p(\boldsymbol{\theta} \mid \boldsymbol{x}_{1:n}, y_{1:n})$   
=  $\arg \min_{\boldsymbol{\theta}} - \log p(y_{1:n} \mid \boldsymbol{x}_{1:n}, \boldsymbol{\theta}) - \log p(\boldsymbol{\theta})$ 

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**Laplace's method** (*P.S. Laplace, 1774*): Let  $\psi(\theta) \triangleq \log p(\theta | \mathbf{x}_{1:n}, \mathbf{y}_{1:n})$ . Take the 2<sup>nd</sup>-order Taylor expansion of  $\psi(\theta)$  around the MAP estimate  $\hat{\theta}$ .

$$\psi(\boldsymbol{ heta}) pprox \psi(\hat{\boldsymbol{ heta}}) + (\boldsymbol{ heta} - \hat{\boldsymbol{ heta}}) \Big[ \dots \Big] + rac{1}{2} (\boldsymbol{ heta} - \hat{\boldsymbol{ heta}})^{ op} \Big[ 
abla^2_{\boldsymbol{ heta}} \psi(\boldsymbol{ heta}) \Big]_{\hat{\boldsymbol{ heta}}} (\boldsymbol{ heta} - \hat{\boldsymbol{ heta}})$$

Compare this to the log-p.d.f. of a multivariate Normal distribution (MVN).

$$\log \mathcal{N}(oldsymbol{ heta};oldsymbol{\mu},oldsymbol{\Sigma}) = -rac{1}{2}(oldsymbol{ heta}-oldsymbol{\mu})^{ op} \, oldsymbol{\Sigma}^{-1} \, (oldsymbol{ heta}-oldsymbol{\mu}) + ext{const.}$$

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The Laplace approximation q of p is a MVN (mean  $\hat{\theta}$ , cov. matrix  $\hat{\Sigma}$ ).

$$q(oldsymbol{ heta}) riangleq \mathcal{N}(oldsymbol{ heta}; oldsymbol{\hat{ heta}}, oldsymbol{\hat{\Sigma}}) \quad oldsymbol{\hat{ heta}} = rg\max_{oldsymbol{ heta}} rac{\psi(oldsymbol{ heta})}{oldsymbol{\Sigma}} \quad oldsymbol{\hat{\Sigma}} = -\left[
abla_{oldsymbol{ heta}}^2 \psi(oldsymbol{ heta})
ight]_{oldsymbol{ heta}}^{-1}$$

- Beware!  $\hat{\Sigma}$  needs to be symmetric and positive semi-definite.
- However  $\psi(\theta)$  being twice continuously diff. around  $\hat{\theta}$  is sufficient:
  - Symmetric: Order of differentiation does not matter.
  - Positive semi-definite: Hessian taken at  $\hat{ heta}$  is negative semi-definite.

#### Laplace approximation: Predictive posterior

How do we make predictions using our Laplace approximation?

$$p(y^* | \mathbf{x}^*, \mathbf{x}_{1:n}, y_{1:n}) = \int p(y^* | \mathbf{x}^*, \theta) p(\theta | \mathbf{x}_{1:n}, y_{1:n}) d\theta$$
$$\approx \int p(y^* | \mathbf{x}^*, \theta) q(\theta) d\theta$$
$$= \mathop{\mathbb{E}}_{\theta \sim q} \left[ p(y^* | \mathbf{x}^*, \theta) \right]$$





(a) Step 1: Find MAP (b) Step 2: Fit approx. (c) Step 3: Predict! Figure: Fig. 1 of (*Daxberger et al., 2021*)

LA hinges on inverting Hessian of the log-posterior distribution w.r.t. heta

$$\hat{\boldsymbol{\Sigma}} = -\left[\nabla_{\boldsymbol{\theta}}^2 \psi(\boldsymbol{\theta})\right]_{\hat{\boldsymbol{\theta}}}^{-1} = -\left[\nabla_{\boldsymbol{\theta}}^2 \log p(y_{1:n} \,|\, \boldsymbol{x}_{1:n}, \boldsymbol{\theta}) + \nabla_{\boldsymbol{\theta}}^2 \log \frac{p(\boldsymbol{\theta})}{p(\boldsymbol{\theta})}\right]_{\hat{\boldsymbol{\theta}}}^{-1}$$

• Known to work with small neural networks. (D.J.C. MacKay, 1992)

- However the Hessian quickly becomes problematic as P grows.
  - Auto-diff. frameworks help but Hessian still needs  $\mathcal{O}(P^2)$  storage.
  - Inversion can also be difficult. (compute & numerical instability)
  - Not guaranteed to be negative semi-definite! (e.g. saddle points)

Fisher information: Avoid 2<sup>nd</sup>-order differentiation and focus on the score.

$$\hat{\boldsymbol{\Sigma}} \approx \boldsymbol{\Sigma}_{\mathsf{Fish.}}, \ \boldsymbol{\Sigma}_{\mathsf{Fish.}} \triangleq -\left[\sum_{i=1}^{n} \sum_{y \sim \boldsymbol{P}(\dots)} \left[ \left[ \nabla_{\boldsymbol{\theta}} \log \boldsymbol{p}(y \mid \boldsymbol{x}_{i}, \boldsymbol{\theta}) \right]_{\hat{\boldsymbol{\theta}}}^{2} \right] + \dots \right]^{-1}$$

Note: Guaranteed positive semi-definite. (cf. info. geometry, natural grad.)

**Gen. Gauss-Newton (GGN):** Let  $\mathcal{J}_{\theta}(\mathbf{x}) \triangleq \left[\nabla_{\theta} f_{\theta}(\mathbf{x})\right]_{\hat{\theta}}$  denote Jacobian.

$$\hat{\boldsymbol{\Sigma}} \approx \boldsymbol{\Sigma}_{\text{GGN}}, \ \boldsymbol{\Sigma}_{\text{GGN}} \triangleq -\left[\sum_{i=1}^{n} \mathcal{J}_{\boldsymbol{\theta}}(\boldsymbol{x}_{i}) \left[\nabla_{f}^{2} \boldsymbol{p}(\boldsymbol{y}_{i} \mid f)\right]_{\hat{f}} \mathcal{J}_{\boldsymbol{\theta}}(\boldsymbol{x}_{i})^{\top} + \ldots\right]^{-1}$$

Note: Equivalent to the Fisher for most common log-likelihoods.

The full covariance matrices  $\Sigma_{\text{Fish.}}$  and  $\Sigma_{\text{GGN}}$  are still quadratic in the number of parameters P, making them difficult to store and difficult to invert.

**Diagonal factorization:** Keep the matrix diagonal and ignore the off-diagonal elements. Lightweight but highest tradeoff in approximation fidelity.

**Low-rank factorization:** Use a low-rank approximation, such as truncated singular value decomposition (SVD). Can be combined with the diagonal factorization.







**Block-diagonal factorization:** Factorize across neural network layers by approximating diagonal blocks for each layer  $\ell$ , i.e. making the assumption of layer independence.



Kronecker factorization (*Ritter et al., 2018*): The approximate blocks for a single data point can be formulated as a Kronecker product (K-FAC).

$$-\left[\mathbf{\Sigma}_{\mathsf{GGN}}^{(\ell)}\right]^{-1} = \sum_{i=1}^{n} \left[\mathbf{A}_{i}^{(\ell)} \otimes \mathbf{B}_{i}^{(\ell)}\right]$$
$$\approx \left[\sum_{i=1}^{n} \mathbf{A}_{i}^{(\ell)}\right] \otimes \left[\sum_{i=1}^{n} \mathbf{B}_{i}^{(\ell)}\right] \triangleq \mathbf{A}^{(\ell)} \otimes \mathbf{B}^{(\ell)}$$

Note:  $\mathbf{A}^{(\ell)} \in \mathbb{R}^{D_{out}^{(\ell)} \times D_{out}^{(\ell)}}, \mathbf{B}^{(\ell)} \in \mathbb{R}^{D_{in}^{(\ell)} \times D_{in}^{(\ell)}}$  and both are pos. semi. definite.

# Hessian approximation: Predictive misspecification

Misspecification? Used the GGN but now the predictive is underfitting?



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#### Linearized Laplace: Linearization of $f_{\theta}$

Recall that  $\hat{\Sigma}$  depends on the Hessian of the log-**posterior dist.** w.r.t.  $\theta$ .

$$\hat{\boldsymbol{\Sigma}} = -\left[\nabla_{\boldsymbol{\theta}}^2 \log \, \boldsymbol{p}(\boldsymbol{y}_{1:n} \,|\, \boldsymbol{x}_{1:n}, \boldsymbol{\theta}) + \nabla_{\boldsymbol{\theta}}^2 \log \, \boldsymbol{p}(\boldsymbol{\theta})\right]_{\hat{\boldsymbol{\theta}}}^{-1}$$

Focus on the log-likelihood. Let  $\mathcal{H}_{\theta}(\mathbf{x}) \triangleq \left[ \nabla_{\theta}^2 f_{\theta}(\mathbf{x}) \right]$  denote Hessian.

$$\nabla_{\theta} \log p(y_i | \mathbf{x}_i, \theta) = \mathcal{J}_{\theta}(\mathbf{x}_i)^{\top} \left[ \nabla_f \log p(y_i | f) \right]$$
$$\nabla_{\theta}^2 \log p(y_i | \mathbf{x}_i, \theta) = \mathcal{H}_{\theta}(\mathbf{x}_i)^{\top} \left[ \nabla_f \log p(y_i | f) \right]$$
$$-\mathcal{J}_{\theta}(\mathbf{x}_i) \left[ \nabla_f^2 p(y_i | f) \right]_{\hat{f}} \mathcal{J}_{\theta}(\mathbf{x}_i)^{\top}$$

GGN approximation!

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GGN approx. makes the assumption that  $\mathcal{H}_{\theta}(\mathbf{x}_i)^{\top} \left[ \nabla_f \log p(y_i | f) \right]$  is zero.

**Condition 1:** The *residual*  $\left[\nabla_f \log p(y_i | f)\right]$  vanishes for all data points. This is not desired, as it would indicate overfitting, nor is it very realistic.

**Condition 2:** The Hessian  $\mathcal{H}_{\theta}(\mathbf{x}_i)$  vanishes for all input points. This is true for linear networks, and we can enforce it by linearizing our network!

**Local linearization of**  $f_{\theta}$  (*Immer et al., 2021a*): Applying GGN approx. to the Hessian of the likelihood turns underlying model from BNN to GLM.

$$f^{\mathsf{lin.}}_{oldsymbol{ heta}}(oldsymbol{x}) = f_{oldsymbol{ heta}}(oldsymbol{x}) + (oldsymbol{ heta} - oldsymbol{\hat{ heta}})\mathcal{J}_{oldsymbol{\hat{ heta}}}(oldsymbol{x})$$

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# Linearized Laplace: GP formulation

**Function-space equivalence** (*Immer et al., 2021a*): This linearized model in weight-space is equivalent to a Gaussian process (GP) in function-space.

$$\begin{split} m(\boldsymbol{x}) &= \mathop{\mathbb{E}}_{\boldsymbol{\theta}} \left[ f_{\boldsymbol{\theta}}^{\text{lin.}}(\boldsymbol{x}) \right] = f_{\hat{\boldsymbol{\theta}}}^{\text{lin.}}(\boldsymbol{x}) \\ k(\boldsymbol{x}, \boldsymbol{x}') &= \operatorname{Cov}_{\boldsymbol{\theta}} \left[ f_{\boldsymbol{\theta}}^{\text{lin.}}(\boldsymbol{x}), f_{\boldsymbol{\theta}}^{\text{lin.}}(\boldsymbol{x}') \right] \\ &= \mathcal{J}_{\hat{\boldsymbol{\theta}}}(\boldsymbol{x}) \boldsymbol{\Sigma}_{\text{GGN}} \mathcal{J}_{\hat{\boldsymbol{\theta}}}(\boldsymbol{x})^{\top} \end{split}$$

**Function-space predictive** (*Immer et al., 2021a*): Leads to a closed form for sampling from the network. Use Monte-Carlo simulation for predictive.

$$p(f^* | \mathbf{x}^*, \mathbf{x}_{1:n}, y_{1:n}) = \mathcal{N}(f^*; f_{\hat{\theta}}(\mathbf{x}^*), \mathcal{J}_{\hat{\theta}}(\mathbf{x}^*) \mathbf{\Sigma}_{\text{GGN}} \mathcal{J}_{\hat{\theta}}(\mathbf{x}^*)^\top)$$
$$p(y^* | \mathbf{x}^*, \mathbf{x}_{1:n}, y_{1:n}) = \mathop{\mathbb{E}}_{f^*} \left[ p(y^* | f^*) \right]$$

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# Linearized Laplace: Demonstration



Figure: Using GLM as the underlying model. (Fig. 3 of Immer et al., 2021a)



Figure: Mean and uncertainty on banana dataset. (Fig. 4 of Immer et al., 2021a)

**Marginal likelihood** (*Immer et al., 2021b*) : Likelihood of model  $\mathcal{M}$ . Overly simple or complex models have low-probability (Occam's razor).

$$\log p(y_{1:n} | \mathbf{x}_{1:n}, \mathcal{M}) = \int \log p(y_{1:n} | \mathbf{x}_{1:n}, \theta) + \log \frac{p(\theta | \mathcal{M})}{p(\theta | \mathcal{M})} d\theta$$
$$\approx \log p(y_{1:n} | \mathbf{x}_{1:n}, \hat{\theta}) + \log \frac{p(\theta | \mathcal{M})}{p(\theta | \mathcal{M})} - \frac{1}{2} \log \left| \frac{1}{2\pi} \mathbf{\Sigma}^{-1} \right|$$

Note: Model  $\mathcal{M}$  can be choice of architecture, hyper-parameters, etc.

#### Linearized Laplace: Model selection



Figure: Marginal likelihood on toy regression. (Fig. 1 of Immer et al., 2021b)



Figure: Marginal likelihood and test accuracy (Fig. 2 of Immer et al., 2021b)

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### Linearized Laplace: Hyperparam. optimization

**Hyperparam. optimization** (*Immer et al., 2021*): Diff. hyperparams  $\mathcal{M}^{\partial}$ . Alternate between optimizing params.  $\theta$  and hyperparams.  $\mathcal{M}^{\partial}$  online.

$$\mathcal{M}_{t+1}^{\partial} \leftarrow \mathcal{M}_{t}^{\partial} + \gamma \nabla_{\mathcal{M}^{\partial}} \log q(y_{1:n} \,|\, \boldsymbol{x}_{1:n}, \mathcal{M})$$

Note: a.k.a. empirical Bayes or type-II maximum likelihood estimation.



Figure: Optimizing marg. lik. in banana dataset. (Fig. 3 of Immer et al., 2021b)

# Linearized Laplace: PyTorch package



Figure: Overview of laplace-torch. (Fig. 2 of Daxberger et al., 2021)

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Focus on tabular data, d columns of numer. or categ. features  $x_1, \ldots, x_d$ .

**Neural additive models** (*Agarwal et al., 2021*): Handle input columns indep. in separate sub-networks to observe the response as we vary inputs.

$$\mathbb{E}[g(y) \mid \boldsymbol{x}] = f_{\boldsymbol{\theta}}(\boldsymbol{x}) \triangleq f_{\boldsymbol{\theta}_1}^{(1)}(x_1) + f_{\boldsymbol{\theta}_2}^{(2)}(x_2) + \dots + f_{\boldsymbol{\theta}_d}^{(d)}(x_d)$$

Laplace-approximated NAMs (*Bouchiat et al., 2024*): Swap the point estimates with Bayesian neural networks and use linearized Laplace approx.

$$\begin{split} f^{\mathsf{lin.}}_{\boldsymbol{\theta}}(\boldsymbol{x}) &\triangleq f^{(1),\,\mathsf{lin.}}_{\boldsymbol{\theta}_1}(x_1) + f^{(2),\,\mathsf{lin.}}_{\boldsymbol{\theta}_2}(x_2) + \dots + f^{(d),\,\mathsf{lin.}}_{\boldsymbol{\theta}_d}(x_d) \\ f^{(j),\,\mathsf{lin.}}_{\boldsymbol{\theta}_j}(x_j) &\triangleq f^{(j)}_{\hat{\boldsymbol{\theta}}_j}(x_j) + (\boldsymbol{\theta}_j - \hat{\boldsymbol{\theta}}_j)\mathcal{J}_{\hat{\boldsymbol{\theta}}_j}(x_j) \end{split}$$

# Laplace-approximated NAMs: Introduction



Figure: Diagram of LA-NAM architecture in (Bouchiat et al., 2024).

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# Laplace-approximated NAMs: Introduction

Laplace-approximated NAMs (*Bouchiat et al., 2024*): The independence of the subnetworks leads to factorized block-diagonal posterior covariance.

$$q(oldsymbol{ heta}) = \mathcal{N}(oldsymbol{ heta}; oldsymbol{\hat{ heta}}, oldsymbol{\Sigma}_{\mathsf{GGN}}), \quad oldsymbol{\Sigma}_{\mathsf{GGN}} = egin{bmatrix} oldsymbol{\Sigma}_{\mathsf{GGN}}^{(1)} & \ldots & oldsymbol{0} \ dots & \ddots & dots \ oldsymbol{0} & \ldots & oldsymbol{\Sigma}_{\mathsf{GGN}}^{(d)} \end{bmatrix}$$

Block factorization leads to factorized marg. lik. (Immer et al., 2023).

$$\begin{split} \log p(y_{1:n} | \mathbf{x}_{1:n}, \mathcal{M}) &\approx \log p(y_{1:n}, \hat{\boldsymbol{\theta}} | \mathbf{x}_{1:n}, \mathcal{M}) + \frac{1}{2} \left| \frac{1}{2\pi} \boldsymbol{\Sigma}_{\mathsf{Full}}^{-1} \right| \\ &\geq \log p(y_{1:n}, \hat{\boldsymbol{\theta}} | \mathbf{x}_{1:n}, \mathcal{M}) + \frac{1}{2} \sum_{j} \left| \frac{1}{2\pi} \boldsymbol{\Sigma}_{\mathsf{GGN}}^{(j)} \right| \end{split}$$

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### Laplace-approximated NAMs: Toy example



Figure: Demonstration on toy dataset. (Fig. 1 of Bouchiat et al., 2024)

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# Laplace-approximated NAMs: MIMIC-III dataset



Figure: Application to MIMIC-III ICU mortality. (Fig. 2 of Bouchiat et al., 2024)

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# Laplace-approximated NAMs: MIMIC-III dataset



(a) Local predictions of NAM. (Fig. 3 of Bouchiat et al., 2024)



(b) Local predictions of LA-NAM. (Fig. 4 of Bouchiat et al., 2024)

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Laplace-approximated NAMs (*Bouchiat et al., 2024*): The subnetworks are not necessarily mutually independent in the true posterior distribution.

Determine mut. inf. using scalar marginal variances  $\sigma_j^2$ ,  $\sigma_{j'}^2$  and covariance  $\sigma_{j,j'}^2$  with a last-layer Laplace approximation on output weights  $\theta_j$  and  $\theta_{j'}$ .

$$\begin{split} \mathsf{I}(\boldsymbol{\theta}_{j};\boldsymbol{\theta}_{j'}) &= \mathsf{H}(\boldsymbol{\theta}_{j}) + \mathsf{H}(\boldsymbol{\theta}_{j'}) - \mathsf{H}(\boldsymbol{\theta}_{j},\boldsymbol{\theta}_{j'}) \\ &\approx \frac{1}{2} \log \left[ \sigma_{j}^{2} \sigma_{j'}^{2} (\sigma_{j}^{2} \sigma_{j'}^{2} - \sigma_{j,j'}^{2})^{-1} \right] \\ &= \frac{1}{2} \log \left[ 1 - \mathsf{Corr}(\theta_{j},\theta_{j'})^{2} \right]^{-1} \end{split}$$

Select top-k, append their  $f_{\theta_{j,j'}}^{(j,j')}(x_j, x_{j'})$  interaction subnetworks, fine-tune.

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# Laplace-approximated NAMs: Feature interaction



Figure: Feature interaction of LA-NAM. (Fig. 5 of Bouchiat et al., 2024)

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